

Causal Inference: A Tutorial

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Causality in Ancient Greek Philosophy

I would rather discover one causal law than be King of Persia.

— Democritus

We have knowledge of a thing only when we have grasped its cause.

— Aristotle, *Posterior Analytics*

We do not have knowledge of a thing until we have grasped its why, that is to say, its cause.

— Aristotle, *Physics*

Questions on Causation

- ▶ Relevant questions about causation:
 - ▶ the philosophical meaningfulness of the notion of causation
 - ▶ deducing the causes of a given effect
 - ▶ understanding the details of causal mechanism
- ▶ Here we focus on **measuring the effects of causes**, where statistics arguably can contribute most
- ▶ Several statistical frameworks
 - ▶ potential outcomes (J Neyman, DB Rubin)
 - ▶ causal diagrams (J Pearl)

Association versus Causation

- ▶ The research questions that motivate most studies in statistics-based sciences are causal in nature.
- ▶ The aim of standard statistical analysis is to infer **associations** among variables
- ▶ Causal analysis goes one step further; its aim is to infer aspects of the *data generating process*
- ▶ In most cases, ***Association does not imply causation***: behind every causal conclusion there must lie some causal assumption that is not testable.

Notations

- ▶ Treatment (e.g. intervention, exposure) W : we will mostly focus on binary treatments
- ▶ Outcome (e.g. disease status) Y
- ▶ Observed covariates or confounders X
- ▶ Unobserved covariates or confounders U
- ▶ Examples of question of interest
 - ▶ Causal effect of exposure on disease
 - ▶ Comparative effectiveness research: whether one drug or medical procedure is better than the other
 - ▶ Program evaluation in economics and policy

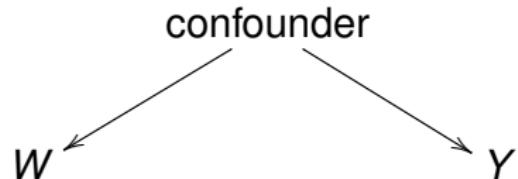
Confounding

- ▶ Confounding (or common cause) is the main complication/hurdle between association and causation
- ▶ Two Directed Acyclic Graphs (Pearl 1995)

Cause relationship:

$$W \longrightarrow Y$$

Confounding:



A Classic Example—Smoking and Lung Cancer

Doll and Hill (1950 BMJ)



Figure: Sir Austin
Bradford Hill
(1897–1991)

- ▶ Smoking-cancer association
- ▶ Case-control study of lung cancer
- ▶ Risk ratio \approx odds ratio, is roughly 9 even after adjusting for observed covariates:

$$RR_{WY}^{\text{obs}} = \frac{\Pr(Y = 1 \mid W = 1)}{\Pr(Y = 1 \mid W = 0)} \approx 9$$

- ▶ Does smoking cause lung cancer?
- ▶ Box (2013) stopped smoking after seeing Doll and Hill (1950)

A Classic Example—Smoking and Lung Cancer

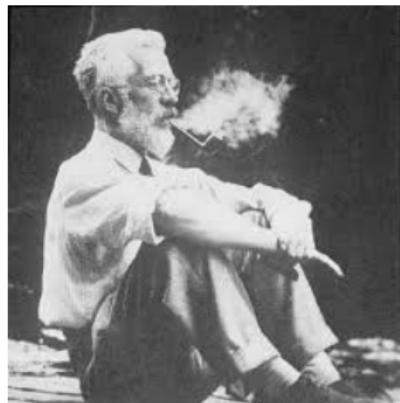


Figure: Sir Ronald Aylmer Fisher (1890–1962)

- ▶ Association does not imply causation
- ▶ “Common cause” (Reichenbach 1956, Fisher 1957 BMJ)
- ▶ Fisher (1957 BMJ):

*cigarette-smoking and lung cancer, though not mutually causative, are both influenced by a **common cause**, in this case the **individual genotype**.*

Simpson's paradox: Kidney Stone Treatment

(Charig et al., BMJ, 1986)

- ▶ An extreme example of confounding is Simpson's paradox: confounder reverses the sign of the correlation between treatment and outcome
- ▶ Compare the success rates of two treatments for kidney stones
- ▶ Treatment A: open surgery; treatment B: small puncture

	Treatment A	Treatment B
Small stones	93% (81/87)	87% (234/270)
Large stones	73% (192/263)	69% (55/80)
Both	78% (273/350)	83% (289/350)

- ▶ What is the confounder here? Severity of the case

Potential Outcome Framework

- ▶ The Potential Outcome Framework: the most widely used framework across many disciplines
- ▶ Brief history
 - ▶ Randomized experiments: Fisher (1918, 1925), Neyman (1923)
 - ▶ Formulation (assignment mechanism and Bayesian model): Rubin (1974, 1977, 1978)
 - ▶ Observational studies and propensity scores: Rosenbaum and Rubin (1983)
 - ▶ Connecting to instrumental variables: Angrist, Imbens and Rubin (1996)

Potential Outcome Framework: Key Components

- ▶ **No causation without manipulation:** a “cause” must be (hypothetically) manipulatable, e.g., intervention, treatment
- ▶ Goal: estimate the **effects of “cause”**, not **causes of effect**
- ▶ Three integral components (Rubin, 1978):
 - ▶ **potential outcomes:** corresponding to the various levels of a treatment
 - ▶ **assignment mechanisms**
 - ▶ a model for the potential outcomes and covariates
- ▶ Causal effects: a comparison of the potential outcomes under treatment and control for *the same set of units*

Setup

- ▶ Data: a random sample of N units from a target population
- ▶ A treatment with two levels: $w = 0, 1$
- ▶ For each unit i , we observe the (binary) treatment status W_i , a vector of covariates X_i , and an outcome Y_i^{obs}
- ▶ For each unit i , two potential outcomes $Y_i(0), Y_i(1)$ – implicitly invoke the **Stable Unit Treatment Value Assumption (SUTVA)**
- ▶ Causal estimands, e.g. **Average treatment effect (ATE)**:
$$\tau = \mathbb{E}[Y_i(1) - Y_i(0)].$$

The Fundamental Problem of Causal Inference

Holland, 1986, JASA

- ▶ For each unit, we can observe at most one of the two potential outcomes, the other is missing (counterfactual?)
- ▶ Causal inference under the potential outcome framework is essentially **a missing data problem**
- ▶ To identify causal effects from observed data, one must make additional (structural or/and stochastic) assumptions
- ▶ Key identifying assumptions are on **assignment mechanism**: the probabilistic rule that decides which unit gets assigned to which treatment

Perfect Doctor

<u>Potential Outcomes</u>		<u>Observed Data</u>		
$Y(0)$	$Y(1)$	W	$Y(0)$	$Y(1)$
13	14	1	?	14
6	0	0	6	?
4	1	0	4	?
5	2	0	5	?
6	3	0	6	?
6	1	0	6	?
8	10	1	?	10
8	9	1	?	9
True averages	7	5	Observed averages	5.4 11

Two key assumptions

Rosenbaum and Rubin, 1983, Biometrika

- ▶ **Strong ignorability** is the key assumption, consisting of
 - ▶ Assumption 1 (**Positivity (a.k.a. overlap)**): each unit has no zero probability of receiving either treatment
 - ▶ Assumption 2 (**Unconfoundedness (a.k.a. ignorability)**): no unmeasured confounders; if two groups have the same distribution of observed covariates, the treatment assignment is random
- ▶ Positivity is testable, but unconfoundedness is generally not

Overlap and Balance

- ▶ Under unconfoundedness, the causal effects are identified from the observed data:
 1. First conditional on subpopulations with covariate balance (via e.g., randomization, or matching, stratification), calculate the difference between treatment and control groups
 2. Average over all such subpopulations (X)
- ▶ The key is to obtain covariate overlap and balance between groups
- ▶ Balance of confounders (observed and unobserved) play a central role in causal inference
- ▶ Observed difference in outcomes might be purely due to the imbalance of confounders between groups

Classification of assignment mechanisms

- ▶ Randomized experiments:
 - ▶ strong ignorability automatically holds
 - ▶ good balance is (in large samples) guaranteed
- ▶ Unconfounded observational studies
 - ▶ strong ignorability is assumed
 - ▶ balance need to be achieved
- ▶ Quasi-experiments: looking for “natural” experiments (under assumptions)

Randomized Experiments

- ▶ In randomized experiments, assignment mechanism is known and controlled by investigators
- ▶ Strong ignorability automatically holds
- ▶ Randomization does:
 - ▶ balance observed covariates
 - ▶ balance unobserved covariates
 - ▶ balance potential outcomes, i.e. guarantee unconfoundedness

Role of Randomization

- ▶ Under randomization, causal effects are identified by the difference in the outcome between the treatment and control groups
- ▶ Under randomization, *association does imply causation* (of course within the potential outcome framework with assumptions)

Chance Imbalance in Randomized Experiments

- ▶ Randomization “should” balance all covariates (observed and unobserved) **on average...**
- ▶ But covariates may be imbalanced by random chance
- ▶ Why is covariate balance important in randomized experiments?
- ▶ Because better balance
 - ▶ Provides more meaningful estimates of the causal effect
 - ▶ **Increases power**, particularly if imbalanced covariates correlated with outcome

Covariate Balance in Randomized Experiments

- ▶ Option 1: force better balance on important covariates *by design* –“Block what you can; randomize what you cannot” (George Box)
 - ▶ stratified randomized experiments
 - ▶ paired randomized experiments
 - ▶ rerandomization
- ▶ Option 2: correct imbalance in covariates by *analysis*
 - ▶ outcome: gain scores
 - ▶ separate analysis within subgroups
 - ▶ covariate adjustment via regression or weighting

Randomized Experiments: Complications

- ▶ Noncompliance
- ▶ Loss to follow-up
- ▶ Truncation due to “death”, e.g. patients died before end of study on life quality
- ▶ Generalize to wider population
- ▶ Ethical and practical constraints: clinical equipoise, sequential trials, pragmatic trials

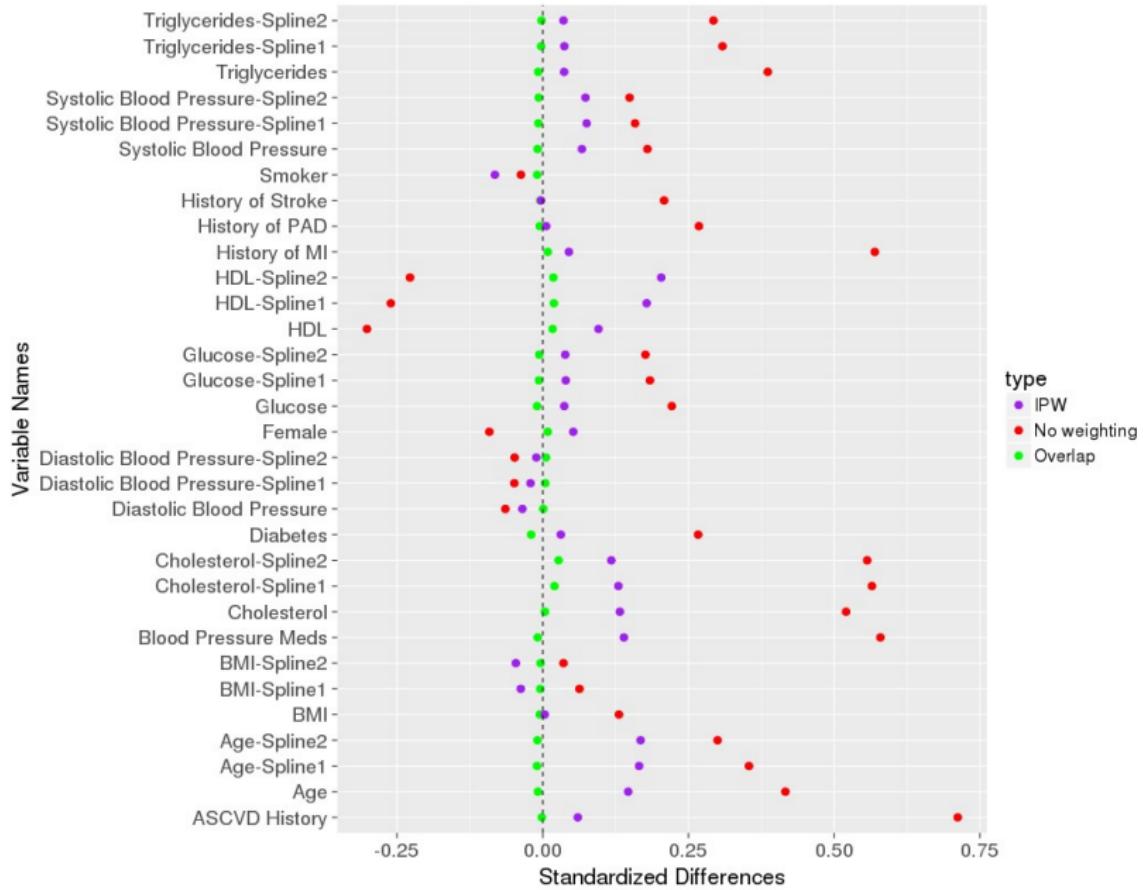
Observational Studies

- ▶ In observational studies, we do not control or know the treatment assignment mechanism
- ▶ Measured and unmeasured confounders: usually unbalanced between groups
- ▶ Self-selection to treatment is prevalent
- ▶ Must make (often untestable) structural assumptions on assignment mechanism to identify causal effects
- ▶ Strong Ignorability is not guaranteed but usually assumed in the vast majority of observational studies

Example: Framingham Heart Study

(Thomas, Lorenzi, et al. 2018)

- ▶ **Goal:** evaluate the effect of statins on health outcomes
- ▶ **Patients:** cross-sectional population from the offspring cohort with a visit 6 (1995-1998)
- ▶ **Treatment:** statin use at visit 6 vs. no statin use
- ▶ **Outcomes:** CV death, myocardial infarction (MI), stroke
- ▶ **Confounders:** sex, age, body mass index, diabetes, history of MI, history of PAD, history of stroke...
- ▶ Significant imbalance between treatment and control groups in covariates



Regression Adjustment

- ▶ Need to adjust difference in the outcomes due to the differences in covariates
- ▶ Most commonly via a regression model:

$$Y \sim a + bW + cX + dW \cdot X$$

- ▶ Potential problems
 - ▶ Regression itself does not take care of lack of overlap or balance
 - ▶ In regions where the groups do not have covariate overlap, causal estimation is purely based on extrapolation
 - ▶ Sensitivity to model-specification

Strategies to Reduce Model Sensitivity

- ▶ To mitigate model dependence, two strategies: (1) design - **balance covariates**, (2) analysis- **flexible models**
- ▶ Best strategy is to use both jointly: first balance covariates in the design stage, then use flexible models in the analysis stage
- ▶ Balance covariates
 - ▶ Stratification or matching
 - ▶ Propensity score methods
- ▶ Flexible models
 - ▶ Semiparametric models (e.g., power series)
 - ▶ Machine learning methods (e.g., tree-based methods (CART, random forest), boosting)
 - ▶ Bayesian non- and semi parametric models (e.g., Gaussian Processes, BART, Dirichlet Processes mixtures)

Balancing covariates: small number of covariates

- ▶ When the number of covariates is small, the adjustment can be achieved by exact matching or stratification
- ▶ Exact matching: for each treated subject, get a control with exact same value of the covariate
- ▶ Exact matching ensures distributions of covariates in treatment and control groups are exactly the same, thus eliminate bias due to difference in X
- ▶ Exact matching is usually infeasible, even with low-dimensional covariates

Matching

- ▶ Regression estimators impute the missing potential outcomes using the estimated regression function
- ▶ Matching estimators also impute the missing potential outcomes, but do so using only the outcomes of nearest neighbours of the opposite treatment group (similar to nonparametric kernel regression methods)
- ▶ Matching is often (but not exclusively) been applied in settings where there is a large reservoir of potential controls

Matching: Dimensional Reduction

- ▶ Matching is good, but...
- ▶ What if there is a large number of covariates? With just 20 binary covariates, there are 2^{20} or about a million covariate patterns
- ▶ Direct matching or stratification is nearly impossible
- ▶ Need dimensional reduction: propensity score

Propensity score

Rosenbaum and Rubin, 1983, Biometrika

- ▶ The propensity score $e(x)$: the probability of a unit receiving a treatment given covariates
- ▶ Two key properties
 1. The propensity score $e(X)$ balances the distribution of all **observed covariates** X between the treatment groups
 2. If the treatment is unconfounded given X , then the treatment is unconfounded given $e(X)$

Propensity score

- ▶ Propensity score is a scalar summary (summary statistic) of the covariates w.r.t. the assignment mechanism
- ▶ Propensity score is central to ensure balance and overlap
- ▶ The propensity score balances the **observed** covariates, but does not generally balance **unobserved** covariates
- ▶ In most observational studies, the propensity score $e(X)$ is unknown and thus needs to be estimated

Propensity score: analysis procedure

Propensity score analysis typically involves two stages:

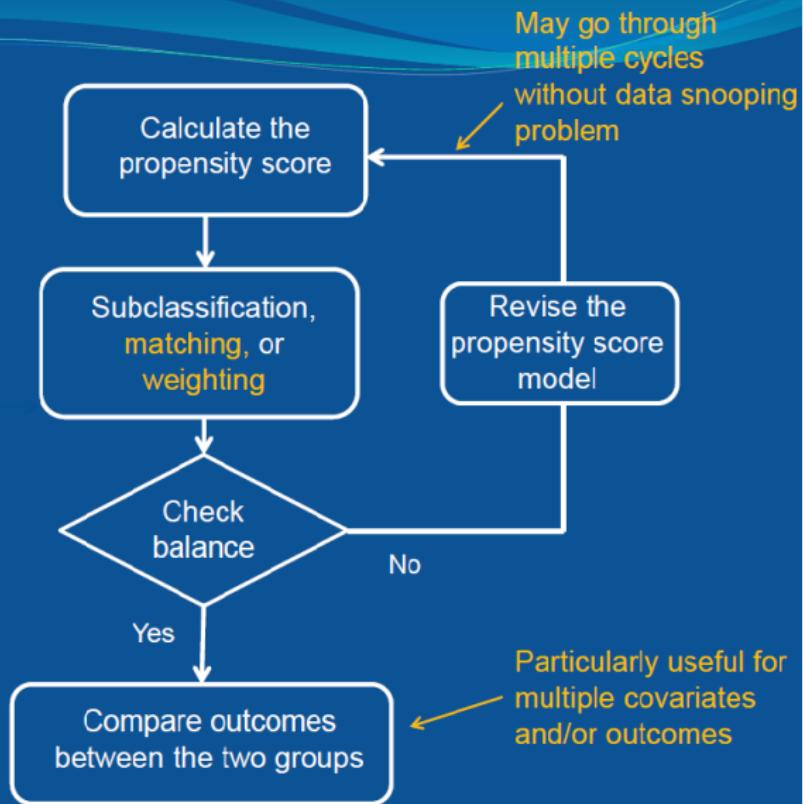
Stage 1 Estimate the propensity score, by e.g. a logistic regression or a machine learning method

Stage 2 Given the estimated propensity score, estimate the causal effects through one of these methods:

- ▶ Stratification
- ▶ Weighting
- ▶ Matching
- ▶ Regression
- ▶ Mixed procedure of the above

Propensity score analysis workflow

propensity score analysis workflow



Propensity score matching

- ▶ Special case of matching: the distance metric is the (estimated) propensity score
- ▶ 1-to-n nearest neighbor matching is common when the control group is large compared to treatment group
- ▶ Pros: intuitive, robust, matched pairs, balance distributions in directions **uncorrelated** to estimated PS
- ▶ Cons
 - ▶ much tuning: with or without replacement, 1-to-1 or 1-to-n, caliper, ties
 - ▶ programming is hard
 - ▶ difficult to extend to complex situations: sequential treatments, multi-valued treatments

Propensity score weighting

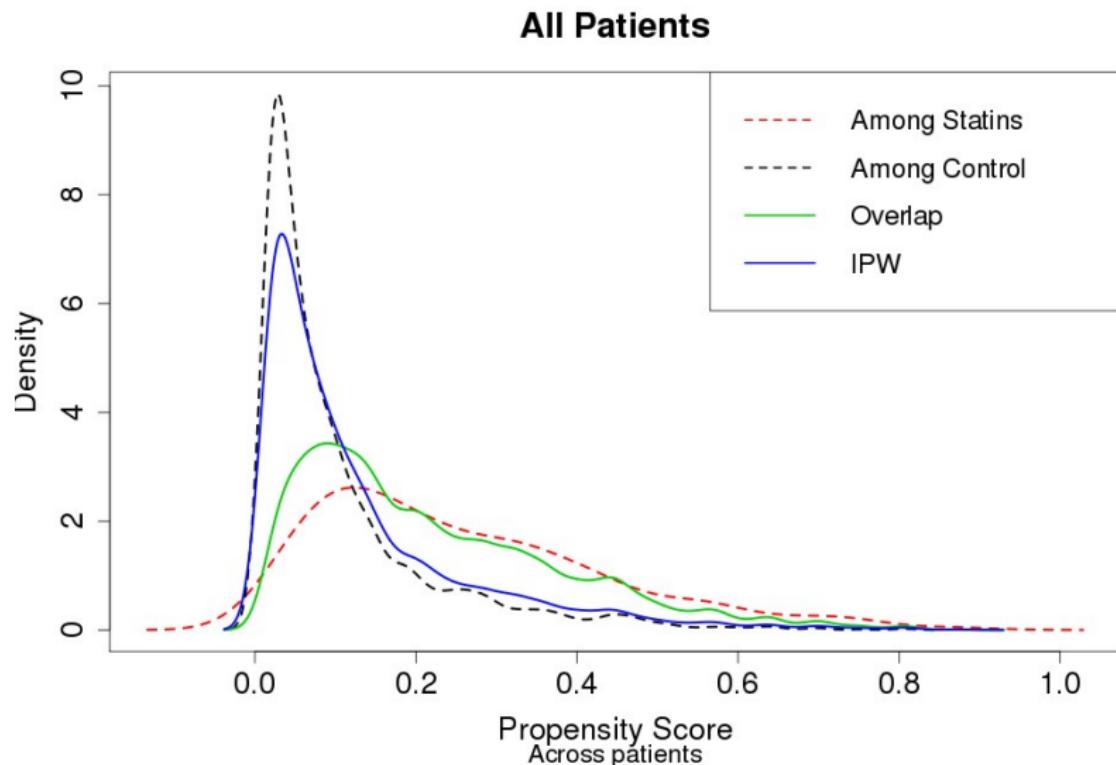
Li, Morgan, Zaslavsky, 2018, JASA

- ▶ Another popular approach is (propensity score) weighting
- ▶ **Main idea:** re-weigh the treatment and control groups to create a pseudo-population—the target population—where the two groups are balanced, in expectation
- ▶ A general class of **balancing weights**
- ▶ Different weighting schemes: different target population and causal estimands
- ▶ One should choose the target population *a priori*

Propensity score weighting: two schemes

- ▶ **Inverse probability weights (IPW)**
 - ▶ Weigh each unit by the inverse of its probability of being assigned to the current group
 - ▶ Target population: the population that the study sample is representative of
 - ▶ But what if the sample is a convenience sample?
- ▶ **Overlap weights** (Li, Morgan, Zaslavsky, 2018, JASA)
 - ▶ Weigh each unit by its probability of being assigned to the **opposite group**
 - ▶ Target population: the population with the most overlap in characteristics between groups (clinical equipoise)
 - ▶ Overlap weights give **exact balance** of covariates

Framingham revisited: weighted distribution



Results: composite of non-death endpoints

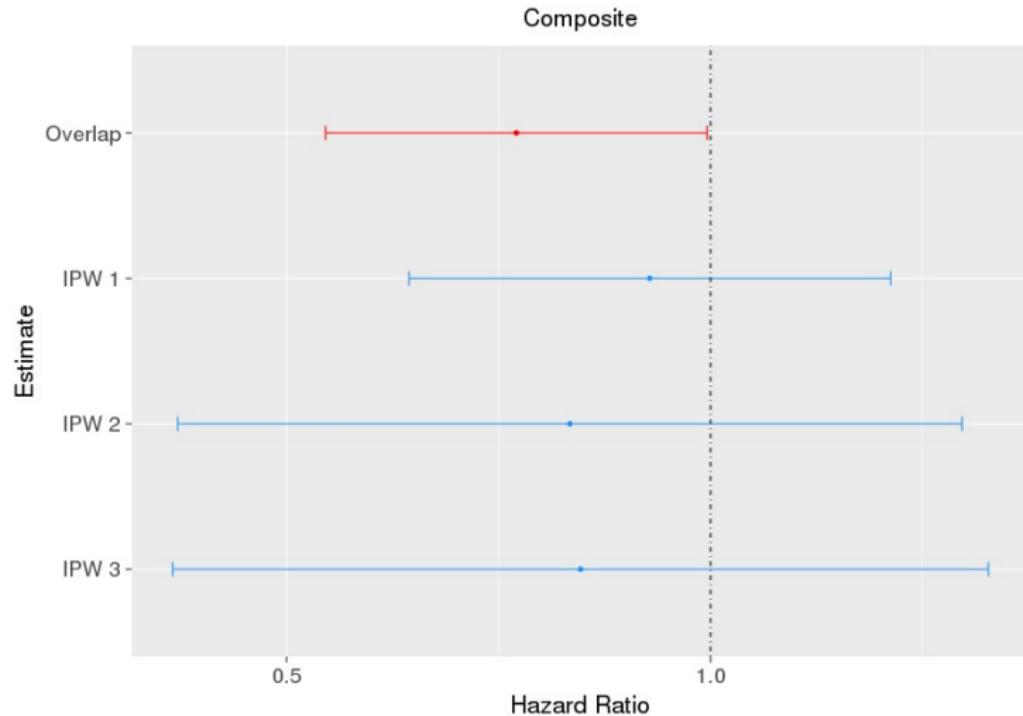


Figure: IPW 1: No trimming; IPW 2: trimming ps between (.10, 0.90); IPW 3: asymmetric trimming 5th% ps of trt, 95th% of ps for control

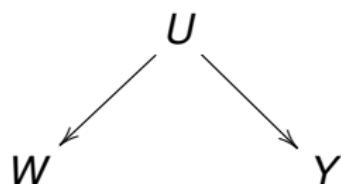
Sensitivity analysis

- ▶ Unconfoundedness is inherently untestable (unknown unknowns)
- ▶ One should always perform sensitivity analysis to assess how sensitive the causal analysis is to violation to unconfoundedness
- ▶ Sensitivity is different from testing, more of a “insurance” check
- ▶ Sensitivity analysis in causal inference dates back to the Hill-Fisher debate on causation between smoking and lung cancer, and first formalized in Cornfield (1959, JNCI)

Smoking and Lung Cancer: Revisited

Cornfield et al., 1959, JNCI

Common cause
hypothesis



- ▶ Smoking W
- ▶ Lung cancer Y
- ▶ Genetic factor U

- ▶ Fisher argued the association between smoking and lung cancer may be due to a **common gene** that causes both
- ▶ Cornfield showed: assuming Fisher is right, the smoking-gene association must satisfy: $RR_{WU} \geq RR_{WY} \approx 9$
- ▶ Such a genetic confounder is too strong to be realistic
- ▶ Thus, here association must be due to causal

Sensitivity analysis

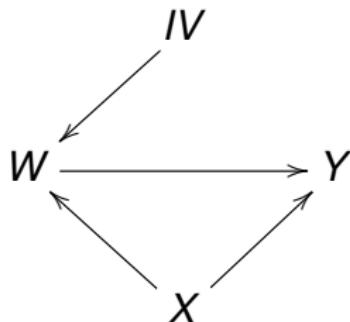
- ▶ Fundamental ideas
 - ▶ Check what would happen to the same analysis had there was an unmeasured confounder? (Rosenbaum and Rubin, 1983b)
 - ▶ Or, how strong an unmeasured confounder has to be to explain away the observed effects? (E-value) (Ding and VanderWeele, 2014, 2016, 2016)
- ▶ Seldom done in substantive field, but should always be checked

Quasi-Experiments

- ▶ Leverage the variation in treatment assignment resulted from nature or policy
- ▶ Three main categories
 - ▶ Instrumental variables (IV)
 - ▶ Regression discontinuity designs (RDD)
 - ▶ Difference-in-Differences (DiD)

Instrumental Variables

- ▶ An instrumental variable (IV): a variable that has a causal effect on the treatment, but (is assumed to) have no “direct” causal effect on the outcome



IV Example 1: Season of Birth

Angrist and Krueger, 1991, Quarterly Journal of Economics

- ▶ Goal: evaluate the effect of schooling on earnings
- ▶ Challenge: Relationship between year of schooling and earnings is highly confounded by factors like family social-economics status
- ▶ IV: **quarter of the year of birth**
- ▶ **Main idea**
 - ▶ When one was born in the year is largely randomized, **by nature**; it should not affect one's later earnings directly
 - ▶ It does directly affects when a child first attended school, and in combination with the compulsory education requirement, this can create up to one year of difference in schooling

IV Example 2: Distance to Hospitals

McClellan, McNeil, Newhouse, 1994, JAMA

- ▶ Goal: evaluate the effect of intensive treatment of acute MI on mortality
- ▶ Challenge: Relationship between receiving intensive treatment and mortality among AMI patients is highly confounded by factors like case severity
- ▶ IV: **distance to the closest hospital**
- ▶ **Main idea**
 - ▶ Where one lives is largely randomized; it should not directly affect one's survival following AMI
 - ▶ It does directly affect what type of hospitals (high vs. low volume and treatment availability) the patient first went, this in turns affects which treatment the patient received

Two-stage Least Square (TSLS) Estimator

- ▶ Traditional model:

$$Y_i = \beta_0 + \beta_1 W_i + \beta_2 X_i + \epsilon_i.$$

where β_1 is the causal effect

- ▶ Direct OLS estimate of β_1 is biased
- ▶ With IV, we can fit a two-stage least square (2SLS) regression to estimate β_1 :

$$M1 : Y_i = \pi_{10} + \pi_{11} Z_i + \pi_{12} X_i + u_i$$

$$M2 : W_i = \pi_{20} + \pi_{21} Z_i + \pi_{22} X_i + v_i$$

- ▶ The 2SLS estimate of β_1 is a ratio:

$$\hat{\beta}_1^{2SLS} = \hat{\pi}_{11} / \hat{\pi}_{21}$$

Instrumental Variables: Assumptions

- ▶ IVs, when available, are extremely useful tools to draw causal inference
- ▶ But, good IVs are hard to come by
- ▶ A good IV must satisfy two conditions (assumptions)
 1. Have a strong effect on the treatment, o.w. the estimate will have large variance
 2. Not have any direct effect on the outcome, o.w. has the same endogenous problem as the treatment
- ▶ Still one of the most popular causal inference methods

Regression discontinuity design (RDD)

- ▶ Regression discontinuity designs: the treatment status changes **discontinuously** according to some underlying pre-treatment variable – the *running variable*
- ▶ **Basic idea:** comparing units with similar values of the running variable, but different levels of treatment would lead to causal effect of the treatment at the threshold
- ▶ The discontinuity is often created by a pre-fixed, artificial threshold of a policy
- ▶ Treatment among the subjects around the threshold can be viewed as **locally randomized**

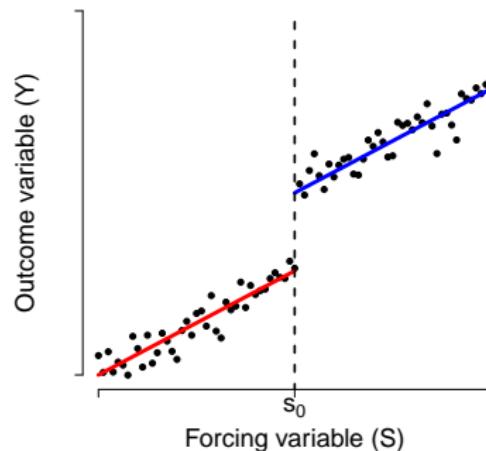
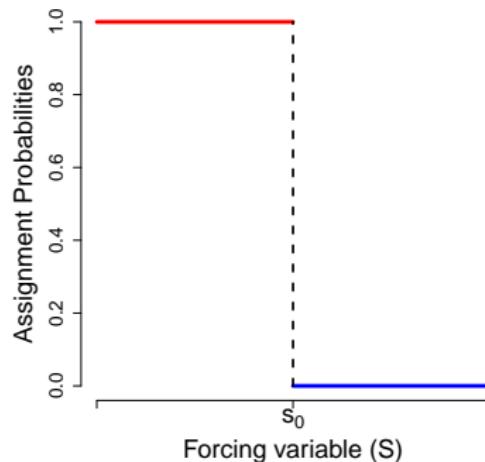
RDD Example: Financial Aid and Dropout

Li, Mattei, Mealli, 2015, AOAS

- ▶ Goal: evaluate the effect of financial aid on preventing dropout in Italian colleges
- ▶ Challenge: students who received aids and who did not are different in observed and unobserved ways
- ▶ Running variable: **family wealth** – eligibility to financial aid depends solely on whether the family wealth is above or below a fixed threshold
- ▶ **Main idea**
 - ▶ Arguably students whose family wealth are just above and just below the threshold are comparable in their background
 - ▶ The artificial threshold set by the administration creates a “local randomization” of treatment around the threshold

Sharp RDD

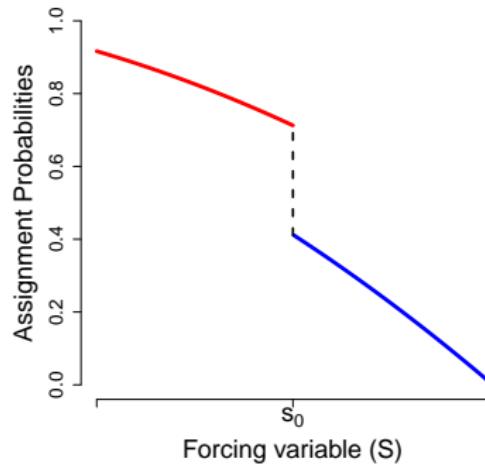
- ▶ The treatment status is a deterministic step function of a running variable



- ▶ Focus on causal effects of the treatment at the threshold
- ▶ A jump in s_0 is interpreted as causal effect

Fuzzy RDD

- ▶ A value of the running variable falling above or below the threshold acts as encouragement to take the treatment



- ▶ In fuzzy RDDs, the receipt of the treatment depends also on individual choices, raising **non-ignorability issues**
- ▶ Fuzzy RDDs are related to IV: falling above or below the threshold can be viewed as an instrument

RDD: assumption and limitationss

- ▶ Key assumption: continuity at the threshold or local randomization
- ▶ Key to analysis: identify a small window around the threshold where local randomization is reasonable
- ▶ Limitations
 - ▶ Treatment effect **local** to the threshold, how generalizable?
 - ▶ Manipulation of the running variable

Difference-in-Differences (DiD)

- ▶ A treatment-control comparison is not necessarily a causal comparison because of the potential systematic differences between two groups
- ▶ A unit is arguably the “best match” for itself
- ▶ A before-after comparison (of the same units) is not necessarily a causal comparison because of the potential change in time
- ▶ Difference-in-Differences (DiD) design combines both: before-after treatment-control comparison
- ▶ Setup: two or more groups, with units observed in two or more periods. In some periods and some groups are exposed to the treatment

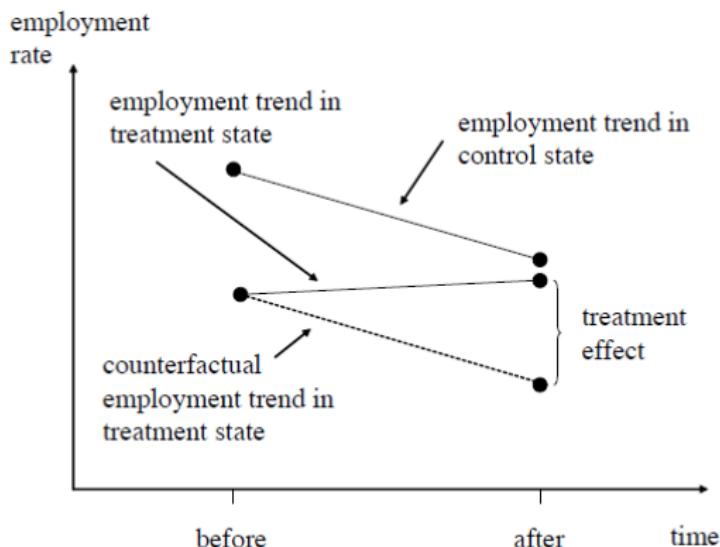
DiD Example: Minimum wages and Employment

Card and Krueger, 1994, American Economic Review

- ▶ Goal: study the effect of increase in minimum wage on employment
- ▶ Units: fast-food restaurants in New Jersey and adjacent eastern PA
- ▶ Intervention: raise of the state minimum wage; NJ raised the minimum on April 1, 1992, but PA not
- ▶ Outcome: number of FTE per restaurant, observed in both areas, and both right-before and after the change
- ▶ **Main idea:**
 - ▶ The restaurants near the state border are arguably similar
 - ▶ Variation in treatment created by discontinuity in time (policy change) and space (state border)

DiD: Parallel Trend Assumption

- ▶ Key assumption: **Parallel trend** – treatment and the control group experience the same trends in the absence of treatment



DiD: Limitations and Alternatives

- ▶ Analysis is usually done via a fixed-effects regression model (time-specific and unit-specific effects)
- ▶ Limitations
 - ▶ Parallel trend is untestable and may be implausible
 - ▶ Scale-dependent: parallel trend for Y does not transfer to $\log Y$
 - ▶ Serial correlation between observations
- ▶ Alternatives: **unconfoundedness** conditional on past outcomes and covariates
- ▶ Uncounfoundedness is also untestable

Final words

- ▶ Causal inference is hard, but fundamental and exciting
- ▶ The potential outcome framework
- ▶ Make and check assumptions
- ▶ Design is the key
- ▶ Many open questions